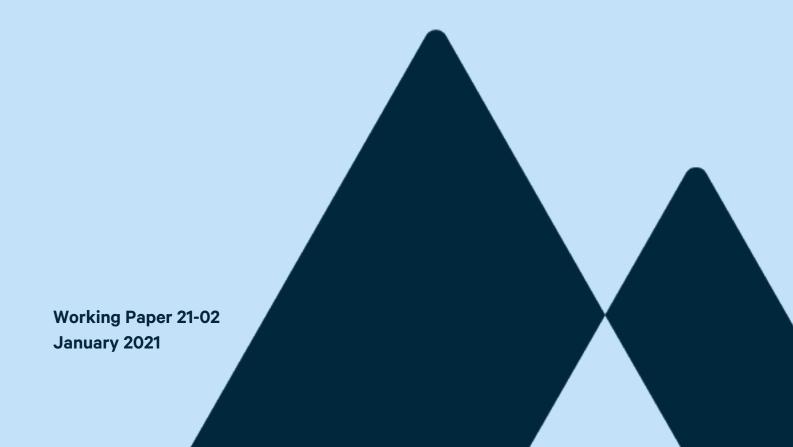


The Climate Decade: Changing Attitudes on Three Continents

Fredrik Carlsson, Mitesh Kataria, Alan Krupnick, Elina Lampi, Åsa Löfgren, Ping Qin, Thomas Sterner, and Xiaojun Yang



About the Authors

Fredrik Carlsson: Department of Economics, University of Gothenburg, Box 640, SE-40530 Gothenburg, Sweden (email: fredrik.carlsson@economics.gu.se)

Mitesh Kataria: Department of Economics, University of Gothenburg, Box 640, SE-40530 Gothenburg, Sweden (email: mitesh.kataria@economics.gu.se)

Alan Krupnick: Resources for the Future, 1616 P St., NW, Washington, DC 20036, USA (email: Krupnick@rff.org)

Elina Lampi: Department of Economics, University of Gothenburg, Box 640, SE-40530 Gothenburg, Sweden (email: elina.lampi@economics.gu.se)

Åsa Löfgren: Department of Economics, University of Gothenburg, P.O. Box 640, S-405 30 Gothenburg, Sweden (email: asa.lofgren@economics.gu.se)

Ping Qin: School of Applied Economics, Renmin University of China, China (email: pingqin@ruc.edu.cn)

Thomas Sterner: Department of Economics, University of Gothenburg, Box 640, SE-40530 Gothenburg, Sweden (email: thomas.sterner@economics.gu.se)

Xiaojun Yang: School of Public Policy and Administration, Xi'an Jiaotong University, China (email: xjyang@mail.xjtu.edu.cn)

Acknowledgments

The authors acknowledge financial support from the Mistra Carbon Exit research program (the Swedish Foundation for Strategic Environmental Research) and Swedish Energy Agency project 46167-1. We have received valuable comments from two anonymous reviewers, participants of the RWI Empirical Environmental Economics Workshop in Essen, and seminar participants at Umeå University and University of Gothenburg.

About RFF

Resources for the Future (RFF) is an independent, nonprofit research institution in Washington, DC. Its mission is to improve environmental, energy, and natural resource decisions through impartial economic research and policy engagement. RFF is committed to being the most widely trusted source of research insights and policy solutions leading to a healthy environment and a thriving economy.

Working papers are research materials circulated by their authors for purposes of information and discussion. They have not necessarily undergone formal peer review. The views expressed here are those of the individual authors and may differ from those of other RFF experts, its officers, or its directors.

Sharing Our Work

Our work is available for sharing and adaptation under an Attribution-NonCommercial-NoDerivatives 4.0 International (CC BY-NC-ND 4.0) license. You can copy and redistribute our material in any medium or format; you must give appropriate credit, provide a link to the license, and indicate if changes were made, and you may not apply additional restrictions. You may do so in any reasonable manner, but not in any way that suggests the licensor endorses you or your use. You may not use the material for commercial purposes. If you remix, transform, or build upon the material, you may not distribute the modified material. For more information, visit https://creativecommons.org/licenses/by-nc-nd/4.0/.

Abstract

Using identical surveys a decade apart, we examine how attitudes and willingness to pay (WTP) for climate policies have changed in the United States, China, and Sweden. All three countries exhibit an increased willingness to pay for climate mitigation. Ten years ago, Sweden had a larger fraction of believers in anthropogenic climate change and a higher WTP for mitigation, but today the national averages are more similar. Although we find convergence in public support for climate policy across countries, there is considerable divergence in both WTP and climate attitudes within countries. Political polarization explains part of this divergence.

Keywords: Climate change, willingness to pay, climate policy attitudes, political polarization, multicountry, China, United States, Sweden

JEL Classification: Q51, Q54

Contents

1. Introduction	1
2. Survey Design	5
3. Results	10
3.1. Respondent Characteristics	10
3.2. Attitudes toward Climate Change and Climate Policies	10
3.3. Willingness to Pay for Reducing Greenhouse Gas Emissions	13
3.4. WTP for Reducing Greenhouse Gas Emissions and the Social Cost of Carbon	14
4. Econometric Analysis	16
4.1. Changes in WTP And Determinants of WTP	16
4.2. Political Polarization and Climate Policy	20
5. Conclusions and Implications for Global Climate Policy	25
6. References	28
Appendix A	32
Online Appendix. Assumptions And Conversion Of Monthly Household Willingness to Pay (WTP) to WTP/Ton CO2	46

1. Introduction

The last decade can surely be called the Climate Decade. High temperatures and carbon dioxide (CO₂) emissions have set new records, but effective political action has proven elusive. Despite strong scientific evidence of the risks of continued warming and substantial progress with renewable energy, little overall progress has been made toward a coherent set of global climate policies. The only signature result has been the 2015 Paris Agreement. While many individual countries and regions have implemented various policies, ranging from energy efficiency standards to carbon pricing, the scope is often too narrow, and policies remain too weak to deliver on the objectives of the Paris Agreement (World Bank, 2019). There are many reasons why policies have or have not been implemented and why there is a lack of stricter global agreements. One important reason is that climate change has the characteristics of an international social dilemma, where countries have to make sacrifices for the collective good (i.e., the global climate system) but, given ample free-riding possibilities, are hesitant to do so. Since mitigating climate change involves economic costs today with mainly collective benefits in the future, political leaders in democratic countries face challenges in legitimizing such costs to their electorates. A vast body of literature explains the difficulty in agreeing on top-down global policies and effort-sharing rules (see, e.g., Keohane, 2011).

Our study examines whether the increasing warning signs over the last decade that the earth is warming are matched by citizens' demand for stronger policy action. Specifically, we address three important questions: (i) how much citizens are willing to pay for reducing CO_2 emissions, (ii) whether citizens' willingness to pay (WTP) has changed over the past decade, and (iii) whether political polarization has increased or decreased when it comes to opinions on climate policies and preferences for decreasing CO_2 emissions. We do this in three key geographies: the United States, China, and Europe (represented in this case by Sweden).

To answer these three questions, we estimate WTP across the three continents in both 2009 and 2019, using the same survey design and the same method (contingent valuation). This comparison with consistent methodology provides a unique opportunity to investigate whether attitudes and WTP have changed both within each country and across the three countries over the decade since the Copenhagen Accord. To our knowledge, this kind of comparison over time has not been undertaken before.

In our comparative study from 2009 (Carlsson et al., 2012), we found that Americans and Chinese citizens were willing to give up about the same percentage of their income for reducing carbon emissions (controlling for income differences), while Swedes were willing to give up a larger fraction of their income. A larger share of Americans were skeptical of climate change and the responsibility of human activities, compared with the Chinese and Swedes. Since that study, a number of cross-country studies have been conducted, several of which include comparisons of US and Chinese citizens (see, e.g., Jamelske et al., 2015; Ziegler, 2017; Schwirplies, 2018; Winden et al., 2018).

Attitudinal changes over time could result from many causes, including increasing political polarization. In the United States, the climate debate has long been polarized between Republicans and Democrats, even well before 2009 (Malka et al., 2009), and in Sweden, signs of increased polarization between sympathizers of different political parties regarding environmental and climate issues have emerged over the last decade (Martinsson and Weissenbilder, 2019). Guber (2013) warns that polarization allows political parties to frame the debate over climate change, with troubling implications. Political polarization could lead to political inaction or weak policies and shift the focus from climate to ideology (Hoffman, 2011). Notably, the United States under the Trump administration has withdrawn from the Paris Agreement, although the incoming Biden administration has announced that the country will immediately rejoin. The election outcome may well be one of the most significant factors in addressing global climate change during the coming decade.

While previous literature has investigated the relationship between WTP for climate policies and political affiliation (Ziegler, 2017; Hornsey et al., 2016; Hornsey et al., 2018) and attitudes and beliefs (Schwirplies, 2018), we instead investigate political polarization related to climate policy and whether this relationship has changed during the past decade. The importance of political ideology in affecting willingness to pay for climate change mitigation does not apply only to Americans. As shown by Hornsey et al. (2016) in a meta-analysis of 171 academic studies across 56 nations, factors such as education, gender, subjective knowledge about climate change, and experience of extreme weather events seem to be less important to explain belief in climate change than ideology, worldview, and political orientation. This result seems particularly important in light of the increased political polarization and the rise of populist parties in the United States and Europe over the last decade (Westfall et al., 2015; Muis and Immerzeel, 2017). Kousser and Tranter (2018) argue that in the presence of political polarization, when political leaders disagree on issues such as climate, the electorate also becomes more polarized.

Our results show that 10 years ago, Sweden stood out, with a larger fraction of believers in anthropogenic climate change and a higher WTP for mitigation. Over the last decade, there have been greater increases in both of these measures in China and the United States, so today the national averages are much more similar. These results are in line with the picture given by individual studies conducted in these countries during the last decade. In particular, there seems to be a trend that Chinese citizens have increased their WTP over time. For example, in the 2018 study by Winden and colleagues, Chinese WTP was more than twice that of their US counterparts (in purchasing power parity terms). Notably, the levels of WTP in 2019 in the three different countries compare closely to conventional measures of the social cost of carbon in China and the United States and to current carbon taxes in Sweden. This is further discussed in both the results and the conclusions sections.

We also find that there is a considerable heterogeneity within the countries, with respect to both attitudes and, more importantly, political affiliation. In both the United States and Sweden, we find a left-right difference in WTP. In the United States, the average WTP of self-identified Republicans is significantly lower than that of people affiliated with other parties, especially Democrats. This is clearly in line with the findings of previous studies, such as Ziegler (2017). Schwirplies (2018) investigates China and the United States (as well as Germany) and finds that political attitudes and socioeconomic characteristics are the main sources of crosscountry heterogeneity. Hornsey and colleagues (2018) find that the correlation between climate skepticism and ideology is both stronger and more consistent in the United States than in 24 other countries. In China, those who are members of the Communist Party have a higher WTP, confirming the previous findings by Ziegler (2017).

Finally, the political divide between Republicans and Democrats in terms of climate policy already existed in 2009, and we find sizable but statistically insignificant changes in polarization in the whole sample of US respondents. This is consistent with the studies on climate attitudes in the United States by Jon Krosnick (PPRG, 2018). Krosnick finds that the partisanship gap in climate beliefs between Republicans and Democrats has been essentially constant since 2011, although it has increased by a small percentage from 2009. However, in our study there is some evidence of increased polarization in the United States if we look at the number of respondents who are not willing to pay anything to decrease carbon dioxide emissions. We find that among self-identified Republicans, this share is large and stable over time, while the corresponding shares among Democrats and independents have significantly decreased further during the past decade. This is an indication of increased polarization between Republicans and others within this

dimension (zero WTP). In Sweden, our results provide stronger evidence of increased polarization with respect to WTP for a stricter climate policy. We find an increasing polarization between left-wing and right-wing voters, and the share of right-wing voters with a zero WTP is significantly larger than the corresponding shares of other voters. Still, it should be noted that in both Sweden and the United States, there is still an overall increase in support for climate policy over the last decade.

The rest of the paper is organized as follows: In section II we describe the design of the studies in 2009 and 2019. In section III we present the results, and in section IV we discuss our findings and their implications.

2. Survey Design

To be able to rule out that any changes across time are not due to changes in the survey design, we used, as far as possible, the same survey in both 2009 and 2019. This comes at a potential cost because the standards for survey design have changed during the last 10 years. However, since the focus of the current study is on comparisons across time, this should not be of any major concern. One of the very few differences between the surveys is that the 2009 survey was longer than the 2019 survey. More specifically, the 2009 survey included a section at the end that was not relevant for the current study. Following are descriptions of the three segments of questions that were included in both the 2009 and 2019 surveys.

In the first segment, we asked general questions about climate change to establish the degree to which respondents are climate change deniers. We asked whether respondents believed that the world's average temperature has increased during the past 100 years, and if so, whether this increase has been caused by human behavior. We also asked whether respondents believed that climate change is inevitable or that there are actions that can reduce or stop it.

In the second segment, we provided information about how different levels of global temperature increases would affect environmental outcomes sensitive to climate change, such as harvests, flooding, storms, and threatened ecosystems. This information was based on a set of UN Intergovernmental Panel on Climate Change (IPCC) reports from 2008. Furthermore, we gave the respondents the following information: "There are a small number of experts who disagree with the IPCC consensus, but because the IPCC represents the large majority of climate researchers, this survey will use information from the IPCC." To keep the information as simple as possible and understandable for respondents without statistical literacy, we did not communicate any additional uncertainties. We explained that the temperature increase depends on the quantity of future global CO₂ emissions. To limit the temperature increase to just 2°F, 3°F, or 4°F, yearly CO₂ emissions would need to be decreased from current levels. The information given to the respondents is shown in Table 1. The respondents were given the information that CO2 reductions of 30, 60, and 85 percent correspond to temperature increases of 4°F, 3°F, and 2°F, respectively. In addition, we noted that if the world does not reduce emissions and continues with business as usual, the IPCC experts predict a temperature increase of

¹ The section at the end of the 2009 survey was reported in Carlsson et al. (2013), and analyzed respondents' preferences for how to divide the global costs to decrease CO2 emissions among different countries.

more than 4°F between 2050 and 2100, which would most likely correspond to even larger changes in the global ecosystem, and most countries would be negatively affected.²

Table 1. The Effects of Temperature Increase on Harvests, Increased Flooding and Storms, and Threatened Ecosystems between 2050 and 2100

Global emissions reduction	85% reduction	60% reduction	30% reduction
Temperature increase	2°F increase	3°F increase	4°F increase
Harvest	Harvests in countries near the equator decrease by 4%–6%. Harvests in countries in the Northern Hemisphere increase by 1%–3%.	Harvests in countries near the equator decrease by 10%–12%. Harvests in countries in the Northern Hemisphere are unaffected.	Harvests in countries near the equator decrease by 14%–16%. Harvests in the Northern Hemisphere decrease by 0%–2%.
Increased flooding and storms	Small tropical islands and lowland countries, for example Bangladesh, experience increased flooding and storms.	Additional low-lying areas in the Americas, Asia, and Africa experience increased flooding and storms.	Populous cities face increased flood risks from rivers and ocean storms. Existence of small island countries is threatened.
Threatened ecosystems	Sensitive ecosystems such as coral reefs and the Arctic ecosystem are threatened.	Most coral reefs die. Additional sensitive ecosystems and species around the world are threatened.	Sensitive and less sensitive ecosystems and species around the world are threatened.

 $^{^2}$ The percentages and temperatures were the same in the 2009 and 2019 surveys, but to keep the time frame the same as in the 2009 survey (40 years), we changed the end year of payment from 2050 to 2060 in the 2019 survey.

After presenting this information, the survey included a set of attitudinal questions about how the respondent's own country should decrease CO2 emissions and whether it should reduce these emissions even if other countries do not. The segment ended with a cheap talk script, 3 followed by three WTP questions asking the respondents to state their household's maximum monthly willingness to pay to decrease CO₂ emissions by 30, 60, and 85 percent (in both surveys, the payment time frame was 40 years). The first WTP question asked respondents how much their household would be willing to pay for a 30 percent reduction compared with no reduction at all (business as usual). The second question asked respondents how much more their household would pay for a 60 percent reduction instead of the 30 percent reduction. The final WTP question asked them how much more their household would be willing to pay for an 85 percent reduction compared with the 60 percent reduction.

A payment card with 24 bids was used when eliciting WTP. The highest bid was \$220, and those who were willing to pay more than that could state their WTP in an open-ended payment question. Respondents who stated zero WTP were also asked all three WTP questions. In each WTP question, total monthly and yearly costs were shown to the respondents to make sure that they understood the consequence of their answers. The payment vehicle for obtaining the different reductions was increases in energy costs, such as through higher electricity and gasoline prices (adjusted for inflation). Respondents were also told that the costs of reducing CO₂ emissions would be shared among the countries according to their current emissions. There were several reasons for using the payment card format. First, there was a clear risk of a poor bid-vector design for at least one country if we would use a closed-ended format, given the uncertainties about the underlying true WTP function. Second, respondents could be uncertain about their WTP, and a payment card format that allows for a wider range of response compared with a discrete choice format could therefore be preferred (Welsh and Poe, 1998). Finally, prior

³ The cheap talk script read, "Before making your choices, please consider how an increased cost would affect your possibilities for buying other things. Previous studies of this kind have shown that people claim to be willing to pay more money than they actually would in a real situation. It is important to us that you answer the questions in this study as truthfully as possible." As Johnston and colleagues. (2017) discuss, the use of cheap talk is potentially problematic, particularly if it affects the perceived consequentiality for some respondents. However, since the script was included in the study conducted in 2009, we decided to keep it in the 2019 study as well. It should be noted that respondents were told that the purpose of the study was to design better climate policies. We therefore have no reason to believe that respondents viewed our study as less consequential than other, similar types of studies, but we did not include any follow-up question on perceived consequentiality (Vossler et al., 2012) or vary the message regarding consequentiality of the survey (Meginnis et al., 2018; Zawojska et al., 2019).

experience with administering both closed-ended and payment card questions in China (Krupnick et al., 2010) led us to favor the latter, based on performance of external scope tests.

The third (and final) segment of the survey included background questions for respondents, mainly about socioeconomic characteristics such as gender, age, education, occupation, and income.

The first study was conducted in November–December 2009. The Chinese survey was administered in four cities: Beijing, Nanning, Jiujiang, and Chonqging. Respondents were randomly selected using neighborhood-based databases, and they were invited to take the survey in special rooms with laptops. These cities were chosen by the Chinese government's Ministry of the Environment (now Ministry of Ecology and Environment) as broadly representative of Chinese cities in terms of size, location, and income (Krupnick et al., 2010). The US and Swedish respondents answered the survey online and were recruited through two large online panels. The 2009 survey yielded 1,221 usable responses in Sweden, 989 in the United States, and 1.264 in China.

The second study was conducted as a web survey during February–April 2019 in all three countries. Before the main study, a pilot study was conducted in each country. The 2019 survey yielded 1,480 usable responses in Sweden, 1,436 in the United States, and 1,557 in China. The Swedish and US respondents were recruited from random samples of a representative panel from each population. The Chinese respondents were recruited from a panel that gave the opportunity to sample from the same four cities as in the 2009 survey. The 2019 panel in China had very few older respondents, creating differences between the two waves in China. To be able to compare the 2009 and 2019 waves, we decided to make a round of additional data

⁴ In total, we received 1,497 survey replies from Sweden, 1,461 from the United States, and 1,579 from China in 2019, but some had to be excluded from the study because of missing observations or extreme values of stated WTP (>\$220 per month). The 2009 survey had 1,229 responses in Sweden, 1,000 in the United States, and 1,448 in China.

⁵ The Swedish respondents were representative in terms of age, gender, median income, and geography when compared with numbers at the national level (Statistics Sweden, 2018a, b, c), but university-educated respondents were overrepresented (Statistics Sweden, 2018d). The respondents from the US panel were representative in terms of age (CIA, 2018a), gender (US Census Bureau, 2018a), and geography (US Census Bureau, 2018b), but the share of university-educated respondents was too large compared with numbers at the national level (US Census Bureau, 2019). For China, the share of young people was overrepresented compared with national statistics (CIA, 2018b).

collection in China in which we targeted older respondents. This data collection was done in December 2019.⁶

Participants in China were invited to take the survey in special rooms with laptops to collect data in 2009, whereas a web-based survey was used in 2019, so potential survey mode effects cannot strictly be excluded. The sense of observability and importance could presumably be stronger in the 2009 survey. In an extensive review, however, Lindhjem and Navrud (2011) conclude that the literature reports similar WTP estimates and no substantial differences between internet surveys and other survey modes. Our results are that the mean WTPs are substantially higher in 2019 than in 2009, which goes in the opposite direction from any survey mode effect.

⁶ An additional 405 interviews in total were conducted on December 2–10, 2019, in Nanning (109), Jiujiang (103), Chongqing (103), and Beijing (90) to complement the data with Chinese respondents age 55 and older. Interviewers visited parks, residential areas, shopping centers, and elderly homes. Representativeness was secured by conducting the interviews in different types of areas and having quotas for age and gender.

3. Results

3.1. Respondent Characteristics

The econometric analysis is based on 2,458 observations from the United States, 2,726 observations from Sweden, and 2,843 observations from China. Descriptive statistics for the demographic variables for the three countries in the two waves are shown in Appendix Table A1. Overall, most demographic variables are stable within countries across years, with a few exceptions. For both Sweden and the United States, the 2019 sample includes more university graduates. In addition, in Sweden, political preferences have changed from 2009 to 2019. In particular, support for the conservative and nationalistic right-wing party, the Sweden Democrats, tripled during this period (Valmyndigheten, 2019a, b). For the China sample, the main differences are that monthly income was substantially higher in the 2019 sample and that the average age was only around 40 years in 2019, whereas it was 54 years in 2009. Note also that the share of members of the Communist Party was relatively high in both waves, with a national average of around 6 percent, although this is likely higher in the urban population. In the analysis, we apply survey weights based on the correct level of educational attainment of the population in all three countries and in 2009 and 2019, respectively. For the Chinese sample, this also partly dealt with the overrepresentation of members of the Communist Party, since membership is more prevalent among those with a higher education. In addition, the main analysis comparing differences between the two waves controls for respondent characteristics.

3.2. Attitudes toward Climate Change and Climate Policies

Descriptive statistics of the most important attitudinal variables are presented in Table 2. The mean values in Table 2 are corrected to account for nonrepresentativeness of the samples with respect to educational attainment in all three countries using poststratification weights (see Appendix Table A4). Differences in means are tested using a Wald chi-squared test.

⁷ For the Swedish and US samples, the larger shares of university graduates were mostly due to an overrepresentation of respondents with a university degree in the panels (see Appendix Table A2). For the Chinese sample, the overrepresentation of university graduates in our sample was lower in 2019 than in 2009 (see Appendix Table A3).

Table 2. Attitudes toward Climate Change: Share of Respondents Agreeing with Each Statement

	United States			Sweden		China			
Statement	2009	2019	p-value ^a	2009	2019	p-value ^a	2009	2019	p-valueª
Global temperature has not increased.	24%	16%	0.000	6%	7%	0.446	5%	3%	0.011
Humans have affected the temperature.	73%	84%	0.000	94%	96%	0.042	94%	98%	0.001
We cannot do anything to stop climate change.	17%	13%	0.071	7%	6%	0.158	11%	13%	0.257
We can mitigate but not stop climate change.	59%	66%	0.011	80%	79%	0.366	77%	76%	0.639
We can stop climate change.	11%	13%	0.275	10%	14%	0.018	9%	9%	0.728
The climate is not changing.	12%	7%	0.000	2%	1%	0.060	2%	1%	0.028
Own country should reduce emissions, even if other countries do not.	68%	79%	0.000	81%	73%	0.000	77%	89%	0.000
Own country should primarily use public funding to reduce carbon emissions wherever it is cheapest, even if it means in another country.	16%	26%	0.00	43%	43%	0.966	34%	41%	0.005
Own country should primarily use public funding to reduce carbon emissions in own country.	42%	46%	0.170	40%	36%	0.068	52%	54%	0.346
No opinion or other opinion.	42%	28%	0.000	17%	21%	0.031	14%	5%	0.000
We should deal with climate change primarily by reducing emissions.	59%	67%	0.003	85%	83%	0.259	80%	78%	0.267
We should deal with climate change primarily by adaptation.	15%	18%	0.250	11%	14%	0.044	17%	21%	0.018
Do nothing, because climate change is not a problem.	25%	15%	0.000	4%	3%	0.163	3%	1%	0.002
The information in the survey regarding climate change is trustworthy.	49%	61%	0.000	81%	79%	0.388	86%	92%	0.000
We should prioritize environmental improvements, even if we lose jobs.	40%	60%	0.000	61%	55%	0.013	75%	82%	0.001
No. of observations	1000	1,458		1,229	1,497		1,264	1,579	

Note: Survey weights based on educational attainment. See Appendix Table A4 for details.

In Sweden, attitudes toward climate change were very similar in 2009 and 2019. Most Swedes acknowledge that human activity has caused rising temperatures to some degree and that we can and should mitigate climate change. However, there are a few statistically significant differences, albeit quite small, between 2009 and

 $^{^{\}it a}$ p-value for Wald test of difference in means based on survey weights.

2019. Most of these indicate that Swedish respondents are now somewhat less committed to certain climate change policies. For example, fewer respondents in 2019 thought that Sweden should prioritize the environment even if that means lost jobs (a decrease by 6 percentage points).

Several prominent and significant attitude changes occurred among the US respondents between 2009 and 2019, most of which were likely to make them more positive toward climate change policies. For instance, the share of people who acknowledge an increase in average global temperature rose by 8 percentage points, and the share who believe that humans have contributed to this rose by 11 percentage points. While the share of Swedish citizens who want their own country to reduce CO_2 emissions, even if other countries do not, has significantly decreased during the past 10 years (81 versus 73 percent), the pattern is the opposite in the United States. In 2019, about 79 percent of US citizens in supported this position, while the corresponding number in 2009 was 68 percent. The largest change in the US samples was that in 2019, a significantly larger share (60 versus 40 percent) of US citizens were willing to prioritize the environment even at the expense of lost jobs.

The Chinese respondents also experienced changes in attitudes regarding climate change, most being more positive toward climate policies. For example, in 2019, 89 percent of the Chinese respondents thought that their own country should decrease CO_2 emissions even if other countries do not, increasing from 77 percent in 2009. There was also stronger support for cost-efficient policies and for mitigation policies in 2019.

Our overarching conclusion is that the populations in these countries are converging in their attitudes. The Chinese and Americans are becoming more convinced of the necessity for climate change action and more committed to taking that action.

Comparing across the countries, in general it seems that the Chinese to a greater extent believe that the climate is changing, followed by the Swedes and Americans.

The Chinese also seemed more favorable in 2019 toward their country acting independently of what other countries do to reduce carbon emissions (89 percent), compared with the United States (78 percent) and Sweden (73 percent). In contrast, in 2009, Swedes were the most positive among the three countries about acting independently of other countries. However, the largest changes in climate-related attitudes have occurred in the United States. In the 2009 sample, 24 percent of US respondents did not believe that the global average temperature had increased, but in 2019, that share was 16 percent.

3.3. Willingness to Pay for Reducing Greenhouse Gas Emissions

Table 3 presents household monthly WTP to reduce CO₂ emissions by 30, 60, and 85 percent. We use the number each respondent chose from a payment card matrix as an estimate of their maximum WTP. The WTPs are weighted to adjust for unrepresentativeness of educational attainment in all three countries (see Appendix Table A4 for the poststratification weights). Respondents' WTP is censored at the highest bid on the payment card.⁸ For the United States, 1.5 percent of the respondents exceeded this limit, and for China and Sweden, the corresponding fractions were 0.8 and 0.9 percent, respectively.

Table 3. Attitudes toward Climate Change: Share of Respondents Agreeing with Each Statement

		United States			Sweden			China	
	Mean	Share zero	Share inc.	Mean	Share zero	Share inc.	Mean	Share zero	Share inc.
					2019 wave				
30% reduction	32 (53)	0.21	0.006	29 (42)	0.14	0.008	24 (40)	0.12	0.009
60% reduction	56 (93)	0.19	0.011	52 (78)	0.12	0.013	49 (78)	0.10	0.018
85% reduction	78 (131)	0.18	0.015	71 (110)	0.12	0.019	75 (118)	0.09	0.027
					2009 wave				
30% reduction	20 (42)	0.29	0.005	23 (35)	0.10	0.006	6 (9)	0.21	0.003
60% reduction	32 (67)	0.26	0.007	42 (63)	0.08	0.011	10 (16)	0.18	0.006
85% reduction	42 (88)	0.24	0.010	57 (90)	0.08	0.015	13 (23)	0.17	0.008

Note: Survey weights based on educational attainment (see Appendix Table A4). Standard deviation in parentheses.

The WTP estimates increase with increasing CO₂ reduction level in all three countries. In general, the WTP estimates are significantly higher for the 2019 wave

 $^{^8}$ We adjusted the WTP for individuals who stated a WTP outside the range of the payment card—that is, those who had a WTP larger than the highest bid of \$220 per month (in the US version)—and set it to \$220. This was because these individuals represent only a small proportion ($\approx 1\%$) of the population that potentially can have a great influence on the mean WTP. Mean WTP would be substantially lower if we instead exclude these responses, in particular for the United States. However, comparisons across time and between countries still reveal the same pattern, and we would reach the same conclusions as we do with the full set of observations.

for all three countries and quite similar in absolute values. The mean WTP for the 85 percent reduction levels in 2019 are strikingly similar, at \$78 in the United States, about \$71 in Sweden, and \$75 in China. For the 85 percent reduction level and from 2009 to 2019, China had the largest increase (477 percent), followed by the United States (86 percent). The increase in Sweden was relatively modest (25 percent). Moreover, only 9 percent of the respondents in China were not willing to pay anything to reduce emissions, compared with 12 percent in Sweden and 18 percent in the United States. This statistic has risen since 2009 by 4 percentage points (from 8 percent) in Sweden, but it has fallen by 6 percentage points (from 24 percent) in the United States and by 8 percentage points (from 17 percent) in China. Thus, the decrease in the fraction of zero WTP responses in the United States and China can partly explain the increase in their mean WTP, and it is another example of the convergence across countries we noted in the previous section. One possible explanation for the decrease in the share of zero WTPs in both China and the United States is the large and significant change in attitudes toward climate change and climate policies that we (and others) have documented.

Finally, we also estimate WTP as a share of mean income in each country. We find that Swedish respondents were willing to pay the largest share of their income in 2009, and the shares were very similar in the United States and China. In 2019, WTP measured as a share of income was quite similar across the countries for the 30 percent reduction level. However, at the 85 percent reduction level, the Chinese respondents stand out as willing to pay considerably more, the US respondents had the lowest value, and the Swedish respondents were intermediate.

3.4. WTP for Reducing Greenhouse Gas Emissions and the Social Cost of Carbon

We now compare our WTP estimates with the social cost of carbon (SCC). The SCC is the damage from emitting an additional ton of carbon dioxide and is used as a proxy for the value of reducing carbon emissions in the economics literature (see, e.g., Greenstone et al., 2013; Pizer et al., 2014; Drupp et al., 2018) as well as in policymaking, particularly in the United States (see, e.g., NASEM, 2017; OECD, 2018). Using public data on emissions in combination with population and household statistics, we perform a simple calculation of willingness to pay per ton of reduced carbon emissions based on our 30 percent reduction estimates as presented in

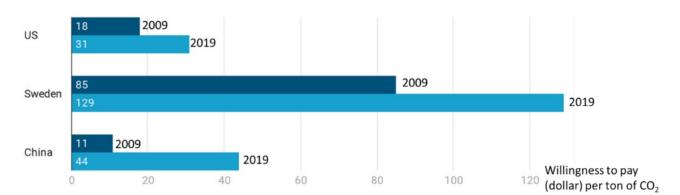


Figure 1. Willingness to Pay per Ton of CO₂ by Country and Year (in dollars)

Table 3.9 These values can then be compared with the range of SCC values usually referred to in the literature (see Rennert and Kingdon, 2019 for an overview of how to calculate the SCC). The WTP estimates for reducing a ton of CO₂ derived from our estimates (for 2009 and 2019, respectively) are illustrated in Figure 1.

Even if our calculations are rudimentary, it is interesting to note that the WTP per ton of reduced CO₂ for each of the three countries in 2019 compares, broadly speaking, with the conventional SCC estimates, even if Sweden is clearly in the higher range. On average, in 2019, Americans were willing to pay \$31/ton CO₂, the Chinese \$44/ton CO₂, and Swedes \$129/ton CO₂. While the US value of \$31/ton CO₂ is the lowest, it is still in the range of values of SCC discussed by the US Environmental Protection Agency (EPA). For Sweden, the \$129/ton CO₂ corresponds almost exactly to the Swedish carbon tax, which is currently around \$126/ton CO₂. In 2009, the average WTP values were considerably lower, and hence the average values attributed to reducing a ton of CO₂ have increased significantly in all three countries. These values have increased fourfold in China and by more than 50 percent in the United States and Sweden.

⁹ See online supplementary material for details on the calculations and data sources used. ¹⁰ EPA's central SCC estimate was \$42/ton in 2017 with a 3 percent discount rate (ranging from \$12 to \$62 for a 5 percent and 2.5 percent discount rate, respectively).

https://19january2017snapshot.epa.gov/climatechange/social-cost-carbon_.html. The Trump administration ordered EPA to calculate a much lower domestic SCC that included damages from climate change only in the United States.

¹¹ The Swedish carbon tax was set to 1,180 SEK/ton CO2 in 2019, which at the exchange rate of 1 US\$ = 9.37 SEK (https://www.forex.se/valuta/usd, accessed November 29, 2020) is 126 US\$/ton CO₂.

4. Econometric Analysis

To investigate the determinants of WTP and its change over the decade, we pool the data for the two waves and estimate a model for each of the three countries separately, using an interval regression model with survey weights to correct for overrepresentation in educational attainment. The interval model was chosen due to the payment card structure, which implies that we only know the ordered categories of WTP but not the exact value of the observation. A dummy variable is included to indicate the 2019 wave. The dependent variable is the interval for the stated WTP in PPP-adjusted dollars at 2018 prices. For brevity, we focus on WTP for reducing emissions by 30 percent compared with the business-as-usual scenario.¹²

4.1. Changes in WTP and Determinants of WTP

For each country, we estimate five models. The first, and simplest, model includes only a dummy variable for the 2019 wave as an explanatory variable (model 1). The model is then expanded with an increasing number of explanatory variables, adding socioeconomic characteristics (model 2), political attitudes (model 3), attitudes toward climate change and policy (model 4), and interaction terms with the dummy variable for the 2019 wave and all other explanatory variables (model 5).¹³

Full results for models 1–5 for a 30 percent reduction are presented in Appendix Tables A5–A7. In this section, we focus on the changes in WTP between the two survey rounds. Table 4 presents the estimates of the difference in WTP between the two survey waves for models 1–4 for the three countries.

¹² Results for 60 and 85 percent reductions for models 1 to 4 are presented in Appendix Tables A8–A10. Results are to a large extent unchanged with respect to statistical significance.

¹³ Models 2–5 contain a set of location control variables. For the United States, we include four regional dummies, with the Southwest region as the reference region. For China, we include city dummy variables, with Beijing as the reference city. For Sweden, we include dummy variables based on the size of the city in which the respondent lives; in total, there are four groups, and the reference category is living in one of the three biggest cities in Sweden. Notably, bigger cities tend to have higher levels of air pollution, and some regions can be more vulnerable to climate change.

Table 4. Estimated Difference between 2009 and 2019 Waves

Sample	Variable	(1)	(2)	(3)	(4)
United States	2019 wave	9.11**	6.85**	7.46***	2.43
		(2.95)	(2.70)	(2.60)	(2.49)
	Pseudo-R ²	0.01	0.09	0.12	0.19
Sweden	2019 wave	4.44***	4.20***	5.13***	6.55***
		(1.25)	(1.23)	(1.25)	(1.24)
	Pseudo-R ²	0.01	0.07	0.10	0.18
China	2019 wave	13.14***	9.02***	9.03***	8.31***
		(0.91)	(1.02)	(1.02)	(1.02)
	Pseudo-R ²	0.08	0.24	0.24	0.25
Location controls		No	Yes	Yes	Yes
Socio-econ controls		No	Yes	Yes	Yes
Political controls		No	No	Yes	Yes
Attitude controls		No	No	No	Yes

Note: Dependent variable is WTP in PPP and inflation-adjusted dollars for reducing emissions by 30% compared with status quo. Standard errors in parentheses. The socioeconomic variables included in the regressions are capturing gender, age, education level, number of adults respectively children in a household, and income level. Variables capturing climate attitudes included opinions as to whether own country should decrease CO_2 emissions regardless of what other countries would do, whether one is for a stricter environmental policy, whether one is for mitigation wherever it is cheapest, and whether one believes that global temperature has not increased. Number of observations are 2,726 for the Swedish sample, 2,458 in the US sample, and 2,843 in the Chinese sample. p < 0.1, p < 0.05, p < 0.01.

Model 1 implies that WTP is \$9.10 higher in 2019 than in 2009 in the US sample, and the corresponding numbers for Sweden and China are \$4.40 and \$13.10, respectively. Adding demographic and political preferences as controls does not significantly affect the difference in mean WTP between 2009 and 2019 in the US sample, but when adding attitudes toward climate change and policies, the 2019 dummy becomes statistically insignificant. Although the pseudo-R2 is low in all the

¹⁴ These estimates are lower than what we saw with the raw data where we used the response on the payment card as the estimate of WTP.

models, there is a sizable increase in the pseudo-R2 when we add the attitudinal variables for the US sample in model 4. For the Swedish sample, the difference increases from \$4.40 to \$5.10 when adding demographic and political preference controls. Moreover, when we add climate change attitudes, the difference between 2009 and 2019 actually increases to \$6.60. Similarly, as for the US sample, adding the attitudinal variables increases pseudo-R2 significantly in model 4. Finally, for China, the difference in WTP decreases to about \$9.00 if we add demographic and party membership controls and to \$8.30 if we add variables capturing attitudes toward climate change and policies. Further, the increase in pseudo-R2 comes mainly from adding demographic variables. For all three countries, a sizable fraction of the increase in WTP is unexplained even when adding a number of control variables.

In model 5, we include interaction terms between the dummy variable for the 2019 wave and all other independent variables. An interaction term reveals whether the increase in WTP from 2009 to 2019 is explained by any of the observable determinants. For example, if the interaction term between the time dummy and the female variable were positive and significant, that would mean that the increase in average WTP between 2009 and 2019 was partly driven by women increasing their WTP more than men. The full model results are presented in Appendix Tables A5–A7, while Table 5 reports only the interaction terms.

There are several significant interaction terms, especially for the United States, where the interaction term for income is positive and statistically significant. This means that the increase in WTP between 2009 and 2019 is explained by higherincome groups increasing their WTP more than other income groups. Moreover, the interaction terms capturing educational levels are all large, negative, and significant, indicating that those with a low education level have increased their WTP relatively more between 2009 and 2019 than those with a higher education level. Moreover, those who believe that the United States should decrease their CO2 emissions regardless of what other countries would do, as well as those who are in favor of a stricter environmental policy, have increased their WTP relatively more than others have during the 10-year period. For Sweden, the increase in the average WTP from 2009 to 2019 was partly driven by the fact that left-wing voters increased their WTP relatively more than right-wing voters. For China, we find that older respondents and those with higher incomes have not increased their WTP as much as younger respondents and those with lower incomes. Moreover, members of the Communist Party and those who are for a stricter environmental policy have increased their WTP relatively more than others.

Table 5. Marginal Effects of Interaction Terms for the United States, Sweden, and China in Model 5

	US	Sweden	China
019 wave	34.68	-14.31*	30.94***
	(22.62)	(8.67)	(4.37)
2019 × Female	-6.10	-2.82	-1.89
	(4.34)	(2.37)	(1.72)
:019 × Age	-0.10	0.12	-0.71***
	(0.13)	(0.08)	(0.07)
2019 × Adults	-3.56	1.84	0.72
	(3.25	(1.93)	(0.93)
019 × Children	-3.76*	0.62	1.47
	(2.20)	(1.44)	(1.73)
2019 × High school	-24.83*	2.36	2.33
-	(14.80)	(3.74)	(1.90)
2019 × Post-high school	-27.72**	2.96	1.83
-	(13.13)	(4.14)	(2.85)
2019 × University	-31.73**	2.57	5.66
•	(14.75)	(4.18)	(3.64)
2019 × Income	0.18***	-0.001	-0.30***
	(0.05)	(0.08)	(0.07)
019 × Democrat	-5.56		,,
	(7.03)		
019 × Other	-10.10		
·- ·- · · · · · · · · · · · · · · · · ·	(6.43)		
019 × Independent	-8.43		
	(5.52)		
019 × Centrist	(0.02)	2.22	
o o ontriet		(4.01)	
019 × Left		6.24**	
Olo Loit		(3.02)	
2019 × Other		6.01	
.010 011101		(3.89)	
2019 × Party member		4.49	7.50***
.o.o raity illeliidei		(3.27)	(2.40)
2019 × Temp. not inc.	0.91	4.49	7.69
.o.a - Temp. not inc.	(3.58)		7.69 (7.76)
2010 × Own country		(3.27)	
2019 × Own country	7.13*	5.33*	2.17
2010 Minimata and afficient	(4.30)	(2.73)	(2.38)
2019 × Mitigate cost-efficient	-2.60	2.05	3.02
	(6.30)	(2.49)	(1.89)
2019 × Stricter env. pol.	8.32*	2.84	5.96***
·	(5.05)	(2.46)	(2.26)

Note: Dependent variable is WTP in PPP and inflation-adjusted dollars for reducing emissions by 30% compared with status quo. Standard errors in parentheses. * p < 0.1, *** p < 0.05, **** p < 0.01. Democrat and independent refer to US party affiliations. In Sweden, "Centrist" consists of the Liberal and Center Parties, and "Left" consists of Left, Social Democrats, and Green Party. "Other" consists of other small parties. The reference group is right-wing parties, such as Republicans in the United States or the Moderate Party, Christian Democrats, and Sweden Democrats in Sweden. In China, the only variable is membership in the Communist Party.

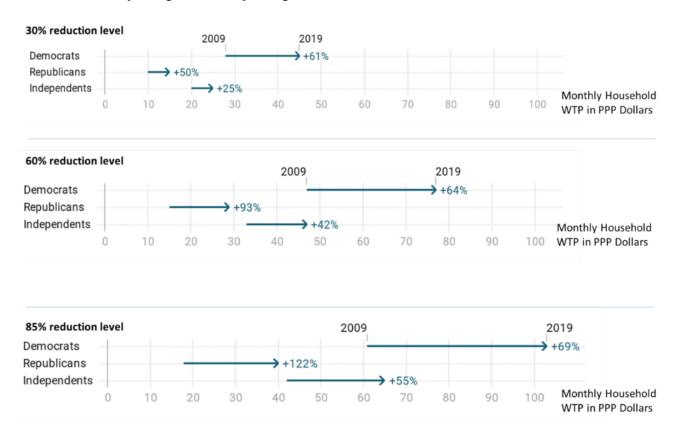
4.2. Political Polarization and Climate Policy

In this section, we explore the issue of political polarization and attitudes toward climate policy in Sweden and the United States. The word polarization is used to describe both a state (a connotation of a division of opinion) and a process whereby opinions become more divided. It is important to note that polarization can be measured in at least two ways: in absolute terms (the difference in WTPs) and in relative terms (the ratio of WTPs). Both measures of polarization can be used to understand the development of climate attitudes and WTP over time.

Relative polarization measures the ratio of WTP of one group compared with another group. Polarization in absolute terms measures the (dollar) difference in WTP of one group compared with another group. Since our focus will mainly be on whether polarization has increased between 2009 and 2019, it is important to note that absolute polarization may increase while the relative polarizations are either unchanged or decreased. For example, assume that WTP is initially twice as high for one group as for the other. Suppose both groups double their WTP at a later point in time. This would result in a doubled difference in absolute WTP but no change in relative WTP between the groups compared across time. Both measures are capturing the extent of polarization but in different units.

Absolute and percentage changes in WTP by (self-reported) political affiliation in the United States are illustrated in Figure 2 (summary statistics, including standard deviations and statistical significance, as well as comparisons of absolute and relative polarization can be found in Appendix Tables A11–A12).

Figure 2. Absolute and Percentage Change in Mean Monthly Household WTP in PPP Dollars (survey weights) Comparing 2009 and 2019 for the United States



Note: The absolute changes are significant at the 1% to 5% significance levels for all the reduction levels for both political parties but not for independents at the 30% level (Wald test, p-value = 0.148).

Given the polarized political discourse during the last decade, exemplified not least by the 2020 US presidential election campaign, one might expect to see increased polarization related to climate change policy between Republicans and Democrats. However, our overall results for the US sample show little support for an increase in polarization; rather, we find a general trend of increased WTP to reduce CO_2 over the last decade among supporters irrespective of party. Comparing absolute changes in WTP, we find a statistically significant increase in polarization only for the 60 percent reduction level, and no statistically significant changes in relative polarization (for details and tests, see Appendix Table A12). Hence, even if there is a clear difference in climate policy opinion between Democrats and Republicans, and the increase in polarization measured as stated WTP between the parties for the last decade is sizable, the increase is not statistically significant. The same pattern holds for Democrats versus independents. The polarization between Republicans and independents in the last decade is negligible.

However, the share of respondents not willing to pay anything for decreasing CO₂ emissions (at the 85 percent reduction level) was significantly higher among Republicans than among both Democrats and independents in 2009 (35 percent versus 12 percent and 25 percent, respectively). While the share of zero WTP has significantly decreased during the past decade among both Democrats and independents, it is about the same among Republicans (see Appendix Table A13). Thus, a stable fraction of over one-third of the Republicans is not willing to pay anything for decreasing CO₂ emissions. In 2009, this share among Republicans was three times larger than that of Democrats, and in 2019, it was more than sevenfold. Hence, there has been an *increasing political polarization* between Republicans and Democrats when it comes to the group of voters who are not willing to pay anything. Put another way, polarization has increased among Republican voters. Similarly, the gap between Republicans and independents not willing to pay anything widened during the past decade.

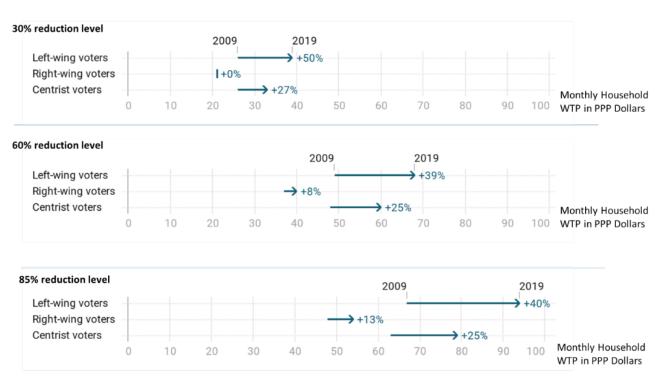
When it comes to climate attitudes, the share of both Democrats and Republicans who believe that humans have affected temperature increases had increased by about 10 percentage points in 2019 compared with 2009 (see Appendix Table A13). Almost all Democratic supporters and around two-thirds of Republican supporters now believe that humans have affected temperature increases. There was also a remarkable increase in support for prioritizing the environment among both Democrats and independents, while the increase in support among Republicans was relatively moderate. However, only among Democratic voters was there a considerable increase (from 80 to 96 percent) among those who think that their own country should decrease its carbon emissions even if other countries do not.

For Sweden, respondents are divided into three main groups: left-wing, centrist, and right-wing voters. Figure 3 illustrates the changes in WTP for the three different political groups (summary statistics, including standard deviations and statistical significance, as well as comparisons of absolute and relative polarization can be found in Appendix Tables A14–A16)

¹⁵ The decrease in the share of zero WTP at the 85 percent reduction level is statistically significant for both Democrats and independents at the 1 percent significance level, while the increase in the share of Republicans with zero WTP is not statistically significant (Wald test, p-value = 0.784).

¹⁶ In Sweden, "Centrist" consists of the Liberal and Center Parties, "Left" consists of the Left Party, Social Democrats, and the Green Party. Right-wing parties consist of the Moderate Party, Christian Democrats, and Sweden Democrats.

Figure 3. Absolute and Percentage Change in Monthly Household WTP in PPP Dollars (survey weights) Comparing 2009 and 2019 for Sweden



Note: According to the Wald test, the absolute changes are significant at all reduction levels for the left-wing voters (p-value = 0.000 for all the levels), and the change at the 30% reduction level is significant for the centrist voters (p-value = 0.083). The changes are all statistically insignificant for the right-wing voters.

Notably, and unlike our findings for the United States, we find evidence of increased polarization in WTP in Sweden for both absolute and relative polarization metrics between left-wing and right-wing voters. However, there was no statistically significant change in polarization between the centrists and either of the other two groups (for details and tests, see Appendix Table A14). Similarly, as is the case with the US sample, the share of right-wing voters with a zero WTP is significantly larger than the corresponding shares of centrists and left-wing voters. Moreover, while the share of those with zero WTP remained stable from 2009 to 2019 among both centrists and left-wing voters, it significantly (p-value = 0.024) increased among the right-wing voters (from 12 to 18 percent).¹⁷

¹⁷ As a robustness test, we reestimate the results in Appendix Table A15, this time excluding the populistic, right-wing conservative party Sweden Democrats from the right-wing parties (see Appendix Table A16). The difference between left- and right-wing voters largely remains, with the exception of the 60 percent reduction level. The share of zero WTP, however, did not significantly change between 2009 and 2019 when we exclude those who sympathize with the Sweden Democrats (Wald test, p-value = 0.258). The share of Sweden

Regarding climate attitudes and beliefs about anthropogenic global warming, there was an increase (from 6 to 11 percent) in right-wing voters who think that we can stop climate change. However, there was a considerable decrease (from 73 to 60 percent) among right-wing voters who think that their own country should decrease its carbon emissions even if other countries do not. In general, less support exists among Swedish right-wing voters for prioritizing the environment, and the share that shows support is considerably lower than among left-wing voters and centrists. Summing up, it is fair to say that political polarization has increased in relation to climate policy in Sweden during the last decade (see Appendix Table A17 for attitudinal results).

Democrat voters has tripled in the country during the time period of our study (2009–19) (Valmyndigheten, 2019 a, b), and the party's support is rising.

5. Conclusions and Implications for Global Climate Policy

The United States and China are the two largest carbon emitters in the world. While US per capita emissions are more than twice as high as in China, the emissions trends over the last decade are quite different. The United States experienced a decreasing emissions trend between 2000 and 2014, while emissions in China increased (World Bank, 2020). Since 2005, China has been emitting more in total than the United States and is the world's largest emitter. Notably, China alone has accounted for two-thirds of global emissions growth over the past decade (IEA, 2018). Europe accounts for less than 10 percent of total global emissions, and Sweden is an insignificant contributor, with per capita emissions that are approximately half of China's and significantly lower than those of the United States.

Despite its marginal contribution to global carbon emissions, Sweden is known for its willingness to take a leading role in climate change mitigation and has adopted an ambitious climate target of zero net emissions by 2045. The United States, on the other hand, has withdrawn from the Paris Agreement under the Trump administration, although the incoming Biden administration has announced that the country will join again. The election outcome may well be one of the most significant factors in addressing global climate change during the coming decade. In addition, the United States has multiple ongoing state initiatives (e.g., California's initiatives for reducing climate change and the Regional Greenhouse Gas Initiative, a cooperative effort among 10 eastern states) and federal regulations (e.g., the Clean Air Act) in place to reduce greenhouse gases. China recently launched a large-scale national power sector emissions trading program to curb carbon emissions and is a leader in renewable energy (Ambrose, 2019), but the country relies heavily on coal, and new investments in coal are being planned at an alarming rate as part of the Belt and Road Initiative (Alvarez, 2019; Saha, 2019). Hence, these three countries reflect very different climate policy contexts, differences that may show up in citizens' attitudes toward climate policies and willingness to pay for climate mitigation.

Interestingly, our results show an increased willingness to pay for climate mitigation in all three countries. Notably, in 2019, the willingness to pay per ton of reduced CO_2 compares closely to conventional measures of the social cost of carbon in China and the United States and to carbon taxes in Sweden. Ten years ago, Sweden stood out with a larger fraction of believers in anthropogenic climate change and a higher WTP for mitigation. Over the last decade, there have been greater increases in both of these measures in China and the United States, so today the national averages are much more similar. Now China stands out with a higher WTP as a share of income.

When interpreting these figures, we need to bear in mind the dramatic changes in these countries over the time period studied. The increased WTP in China should be considered against a background of rapidly rising income, even more rapidly escalating emissions, and a strong increase in stated commitment among political leaders to address climate change. The Swedish numbers may also be seen in the light of a country that has been rather ambitious in its climate efforts but, with a relatively stagnant economy, is now trying to stimulate its economy and coordinate mitigation efforts with other countries in the EU, including countries that are more skeptical toward more stringent climate policies.

What does all this imply for the future of climate policy? A key reason for political inaction could be that public demand for more stringent policies is weak. However, we do not find support for such arguments. Instead, we find that for all three countries, the average WTP per ton of carbon emissions reduced is within the range of the social cost of carbon values usually referred to in the literature. In Sweden, WTP per ton of carbon emissions reduced is comparable to the Swedish carbon tax (approximately \$126/ton), and in the United States, it is below but roughly on par with EPA's central estimates of the social cost of carbon in 2017 (\$42/ton).

On the other hand, within countries, these averages conceal considerable preference divergence, most importantly relating to political affiliation. We find a clear left-right difference in WTP in both the United States and Sweden. In the United States, the average WTP of self-identified Republicans is significantly lower than that of others, especially Democrats. At the same time, this difference in preferences was already present in 2009, and we find little evidence of sizable increases in polarization across Democrats and Republicans when looking at the whole sample. In Sweden, however, which was more homogeneous earlier, there is stronger evidence of increased polarization in both preferences and attitudes toward climate policy. While WTP for climate change policy among right-wing voters has increased, it has increased to a significantly lesser degree than among left-wing and centrist voters. Moreover, in both the United States and Sweden, the share of respondents with zero WTP is significantly higher among Republicans and right-wing party voters than among any other voters, and that gap has increased during the past decade. Hence, while we see convergent tendencies across countries, there is some evidence of increased divergence within countries.

Returning to our initial question whether the increasing warning signs over the last decade that the earth is warming are matched by citizens' demand for stronger policy action, the answer is affirmative and positive for those who have hoped that electorates will demand more stringent climate policies from their leaders. This greater support for climate action could influence future domestic and global policies. However, a consequence of political polarization might be that large groups

of citizens are unhappy with their country's climate politics. If so, political polarization could shift the focus from climate to ideology, and political decisions about climate policies could become more difficult to reach both nationally and globally.

6. References

- Alvarez, C. F. 2019. "Commentary: Fading Fast in the US and Europe, Coal Still Reigns in Asia." December 12. International Energy Agency. https://www.iea.org/commentaries/fading-fast-in-the-us-and-europe-coal-still-reigns-in-asia.
- Ambrose, J. 2019. "China's Appetite for Coal Power Returns despite Climate Pledge." Guardian, November 20. https://www.theguardian.com/world/2019/nov/20/china-appetite-for-coal-power-stations-returns-despite-climate-pledge-capacity.
- Carlsson, F., M. Kataria, A. Krupnick, E. Lampi, Å. Löfgren, P. Qin, S. Chung, and T. Sterner. 2012. "Paying for Mitigation: A Multiple Country Study." *Land Economics* 88 (2): 326–40. https://doi.org/10.3368/le.88.2.326.
- Carlsson, F., M. Kataria, A. Krupnick, E. Lampi, Å. Löfgren, P. Qin, and T. Sterner. 2013. "A Fair Share: Burden-Sharing Preferences in the United States and China." *Resource and Energy Economics* 35 (1): 1–17.
- CIA (Central Intelligence Agency). 2018a. The World Factbook, United States. https://www.cia.gov/library/publications/the-world-factbook/geos/us.html.
- ——. 2018b. The World Factbook, China. https://www.cia.gov/library/publications/the-world-factbook/geos/ch.html.
- Drupp, M. A., M. C. Freeman, B. Groom, and F. Nesje. 2018. "Discounting Disentangled." American Economic Journal: Economic Policy 10 (4): 109–34.
- Greenstone, M., E. Kopits, and A. Wolverton. 2013. "Developing a Social Cost of Carbon for US Regulatory Analysis: A Methodology and Interpretation." *Review of Environmental Economics and Policy* 7 (1): 23–46.
- Guber, D. L. 2013. "A Cooling Climate for Change? Party Polarization and the Politics of Global Warming." *American Behavioral Scientist* 57 (1), 93–115.
- Hoffman, A. J. 2011. "The Growing Climate Divide." Nature Climate Change 1 (4): 195-96.
- Hornsey, M. J., E. A. Harris, P. G. Bain, and K. S. Fielding. 2016. "Meta-Analyses of the Determinants and Outcomes of Belief in Climate Change." *Nature Climate Change* 6 (6): 622–26.
- Hornsey, M. J., E. A. Harris, and K. S. Fielding. 2018. "Relationships among Conspiratorial Beliefs, Conservatism and Climate Scepticism across Nations." *Nature Climate Change* 8 (7): 614–20.
- IEA (International Energy Agency). 2018. CO₂ Emissions from Fuel Combustion 2018. Paris: IEA.
- Jamelske, E., J. Boulter, W. Jang, J. Barrett, L. Miller, and W. L. Han. 2015. "Examining Differences in Public Opinion on Climate Change between College Students in China and the USA." Journal of Environmental Studies and Sciences 5 (2): 87–98. https://doi.org/10.1007/s13412-015-0229-9.

- Johnston, R. J., K. J. Boyle, W. Adamowicz, J. Bennett, R. Brouwer, T. A. Cameron, @. M. Hanemann, et al. 2017. "Contemporary Guidance for Stated Preference Studies." *Journal of the Association of Environmental and Resource Economists* 4 (2): 319–405.
- Keohane, R. O., and D. G. Victor. 2011. "The Regime Complex for Climate Change." Perspectives on Politics 9 (1): 7–23.
- Kousser, T., and B. Tranter. 2018. "The Influence of Political Leaders on Climate Change Attitudes." *Global Environmental Change* 50:100–109.
- Krinsky, I., and L. Robb. "On Approximating the Statistical Properties of Elasticities." Review of Economics and Statistics 68:715–19.
- Krupnick, A., S. Hoffman, and P. Qin. 2010. "The Willingness to Pay for Mortality Risk Reductions in China." Working paper. Washington, DC: Resources for the Future.
- Lindhjem, H., and S. Navrud. 2011. "Using Internet in Stated Preference Surveys: A Review and Comparison of Survey Modes." *International Review of Environmental and Resource Economics* 5 (4): 309–51.
- Malka, A., J. A. Krosnick, and G. Langer. 2009. "The Association of Knowledge with Concern about Global Warming: Trusted Information Sources Shape Public Thinking." *Risk Analysis* 29 (5): 633–47.
- Martinsson, J., and M. Weissenbilder. 2019 "Polariserade problembeskrivningar?" i Andersson m.fl. (red) *Storm och stiltje*. GU: SOM-institutet.
- Meginnis, K., M. Burton, R. Chan, and D. Rigby. 2018. "Strategic Bias in Discrete Choice Experiments." *Journal of Environmental Economics and Management* 102163. https://doi.org/10.1016/j.jeem.2018.08.010.
- Muis, J., and T. Immerzeel. 2017. "Causes and Consequences of the Rise of Populist Radical Right Parties and Movements in Europe." *Current Sociology* 65 (6): 909–30. https://doi.org/10.1177/0011392117717294.
- NASEM (National Academies of Sciences, Engineering, and Medicine). 2017. Valuing Climate Damages: Updating Estimation of the Social Cost of Carbon Dioxide. Washington, DC: National Academies Press. https://doi.org/10.17226/24651.
- OECD (Organisation for Economic Co-operation and Development). 2018. "Current Use of Cost-Benefit Analysis." Chap. 14 in Cost-Benefit Analysis and the Environment: Further Developments and Policy Use. Paris: OECD Publishing. https://doi.org/10.1787/9789264085169-en.
- Pizer, W., M. Adler, J. Aldy, D. Anthoff, M. Cropper, K. Gillingham, M. Greenstone, et al. 2014. "Using and Improving the Social Cost of Carbon." *Science* 346 (6214): 1189–90.
- PPRG (Political Psychology Research Group). 2018. "American Public Opinion on Global Warming." Stanford University. https://climatepublicopinion.stanford.edu/.
- Rennert, K., and C. Kingdon. 2019. "Social Cost of Carbon 101: A Review of the Social Cost of Carbon, from a Basic Definition to the History of Its Use in Policy Analysis." Washington, DC: Resources for the Future.

- Saha, S. 2019. "China's Belt and Road Plan Is Destroying the World." *National Interest*, August 18. https://nationalinterest.org/feature/chinas-belt-and-road-plan-destroying-world-74166.
- Schwirplies, C. 2018. "Citizens' Acceptance of Climate Change Adaptation and Mitigation: A Survey in China, Germany, and the U.S." *Ecological Economics* 145:308–22. https://doi.org/10.1016/J.ECOLECON.2017.11.003.
- Statistics Sweden. 2018a. "Population by Age and Sex: Year 1860–2018."

 http://www.statistikdatabasen.scb.se/pxweb/en/ssd/START_BE_BE0101_BE0101A/BefolkningR1860/ (accessed September 26, 2019).
- ——. 2018b. "Median Salaries in Sweden." [In Swedish.] https://www.scb.se/hitta-statistik/sverige-i-siffror/utbildning-jobb-och-pengar/medianloner-i-sverige/ (accessed September 25, 2019).
- ——. 2018c. "Population by Region, Marital Status, Age and Sex: Year 1968–2018." http://www.statistikdatabasen.scb.se/pxweb/en/ssd/START_BE_BE0101_BE0101A/Be folkningNy/table/tableViewLayout1/ (accessed September 26, 2019).
- ——. 2018d. "Education Level in Sweden." [In Swedish.] https://www.scb.se/hitta-statistik/sverige-i-siffror/utbildning-jobb-och-pengar/utbildningsnivan-i-sverige/?showRelatedArticles=False&showRelatedFacts=True (accessed September 25, 2019).
- US Census Bureau. 2018a. "Quick Facts: Age and Sex." https://www.census.gov/quickfacts/fact/table/U.S./LFE046217 (accessed September 26, 2019).
- ——. 2018b. "Table 1. Annual Estimates of the Resident Population for the United States, Regions, States, and Puerto Rico: April 1, 2010 to July 1, 2018." NST-EST2018-01. https://www.census.gov/newsroom/press-kits/2018/pop-estimates-national-state.html (accessed September 26, 2019).
- ——. 2019. "Table 1. Educational Attainment of the Population 18 Years and Over, by Age, Sex, Race, and Hispanic Origin: 2018; All Races." https://www.census.gov/data/tables/2018/demo/education-attainment/cps-detailed-tables.html (accessed September 26, 2019).
- Valmyndigheten [Swedish Election Authority]. 2019a. "Election Results 2010." [In Swedish.] https://val.se/valresultat/riksdag-landsting-och-kommun/2010/valresultat.html (accessed October 21, 2019).
- ——. 2019b. "Election Results 2018." [In Swedish.] https://www.val.se/servicelankar/other-languages/english-engelska/election-results-2018.html (accessed October 21, 2019).
- Vossler, C. A., M. Doyon, and D. Rondeau. 2012. "Truth in Consequentiality: Theory and Field Evidence on Discrete Choice Experiments." *American Economic Journal:*Microeconomics 4(4): 145–71.
- Welsh, M. P., and G. L. Poe. 1998. "Elicitation Effects in Contingent Valuation: Comparisons to a Multiple Bounded Discrete Choice Approach." *Journal of Environmental Economics and Management* 36(2): 170–85.

- Westfall, J., L. Van Boven, J. R. Chambers, and C. M. Judd. 2015. "Perceiving Political Polarization in the United States: Party Identity Strength and Attitude Extremity Exacerbate the Perceived Partisan Divide." *Perspectives on Psychological Science* 10 (2): 145–58. https://doi.org/10.1177/1745691615569849.
- Winden, M., E. Jamelske, and E. Tvinnereim. 2018. "A Contingent Valuation Study Comparing Citizen's Willingness-to-Pay for Climate Change Mitigation in China and the United States." *Environmental Economics and Policy Studies* 20 (2): 451–75. https://doi.org/10.1007/s10018-017-0202-9.
- World Bank. 2019. "State and Trends of Carbon Pricing 2019." Washington, DC: World Bank Group. https://openknowledge.worldbank.org/handle/10986/31755 (accessed April 28, 2020).
- ——. 2020. "World Development Indicators." Washington, DC: World Bank Group. https://datacatalog.worldbank.org/dataset/world-development-indicators (accessed February 14, 2020).
- Zawojska, E., A. Bartczak, and M. Czajkowski. 2019. "Disentangling the Effects of Policy and Payment Consequentiality and Risk Attitudes on Stated Preferences." *Journal of Environmental Economics and Management* 93:63–84.
- Ziegler, A. 2017. "Political Orientation, Environmental Values, and Climate Change Beliefs and Attitudes: An Empirical Cross Country Analysis." *Energy Economics* 63:144–53. https://doi.org/10.1016/J.ENECO.2017.01.022.

Appendix A

Table A1. Descriptive Statistics of Demographic Variables, Mean Values

Variable	Description	United S	United States		Sweden		China	
		2009	2019	2009	2019	2009	2019	
Female	= 1, if female	0.51	0.55	0.48	0.51	0.46	0.47	
Ago	Years	47.7	49.3	50.7	50.0	54.1	40.0	
Age	redis	(16.84)	(15.7)	(15.4)	(18.2)	(12.6)	(16.4)	
Adults	No. of household members 18 and above	2.18	2.00	1.87	1.79	2.90	2.94	
Adults	No. of flousefiold flieffibers to and above	(0.99)	(0.91)	(0.68)	(0.77)	(1.14)	(1.09)	
Children	No. of household members below 18	0.59	0.64	0.52	0.50	0.52	0.95	
Children	No. of nousehold members below 18	(1.03)	(1.11)	(0.91)	(0.92)	(0.66)	(0.78)	
Compulsory	= 1, if only compulsory/primary school (ref. category)	0.11	0.006	0.11	0.05	0.16	0.20	
High school	= 1, if only high school	0.31	0.10	0.34	0.31	0.13	0.21	
Post-high s.	= 1, if post-high school	0.28	0.31	0.17	0.18	0.28	0.32	
University graduate	= 1, if completed university education	0.30	0.58	0.37	0.46	0.43	0.27	
Income	Annual after-tax income in PPP and inflation-adjusted	49.9	63.8	44.8	45.4	18.8	33.0	
Income	US\$000s	(40.6)	(50.9)	(20.2)	(19.8)	(13.0)	(19.2)	
Miss income	= 1, if no response to income question		0.13					
Political affiliation								
Left	= 1, if left wing (Sweden)			0.43	0.35			
Liberal	= 1, if liberal party			0.13	0.14			
Other party	= 1, if other party (Sweden)			0.107	0.09			
Right	= 1, if conservative right-wing parties (Sweden)			0.33	0.40			
Democrat	= 1, if Democrat (United States)	0.37	0.34					
Other party	= 1, if other party (United States)	0.10	0.09					
Independent	= 1, if independent (United States)	0.28	0.30					
Republican	=1, if Republican (United States; ref. category)	0.25	0.27					
Communist	= 1, if Communist Party (China)					0.31	0.27	
No. observations	•	1000	1,458	1,229	1,497	1,264	1,579	

Note: Standard deviations in brackets.

Table A2. Mean Values Sample Statistics

Variable	Description	United S	United States		Sweden		
		2009	2019	2009	2019	2009	2019
Compulsory	= 1, if only compulsory/primary school	0.11	0.005	0.11	0.05	0.16	0.20
High school	= 1, if only high school	0.31	0.10	0.34	0.31	0.14	0.21
Post-high school	= 1, if post-high school	0.28	0.31	0.18	0.18	0.28	0.32
University	= 1, if completed universityeducation	0.30	0.58	0.37	0.45	0.40	0.26

Table A3. Population Statistics

Variable	Description	United States		Sweden		China	
		2009	2019	2009	2019	2009	2019
Compulsory	= 1, if only compulsory/primary school	0.31	0.38	0.20	0.18	0.27	0.24
High school	= 1, if only high school	0.24	0.23	0.47	0.44	0.39	0.37
Post-high school	= 1, if post-high school	0.31	0.29	0.18	0.15	0.18	0.19
University	= 1, if completed university education	0.14	0.12	0.19	0.23	0.16	0.20

Table A4. Poststratification Weight for Countries

Variable	Description	United States		Sweden		China	
		2.81	7.60	1.81	3.6	1.69	1.2
Compulsory	= 1, if only compulsory/primary school	0.77	2.30	1.38	1.42	2.79	1.76
High school	= 1, if only high school	1.10	0.93	1	0.83	0.64	0.59
Post-high school	= 1, if post-high school	0.47	0.21	0.51	0.51	0.4	0.77
University	= 1, if completed university education	2.81	7.60	1.81	3.6	1.69	1.2

Note: Poststratification weight = population mean / sample mean.

Table A5. Marginal Effects of Interaction Terms for the United States, Sweden, and China in Model 5

	(1)	(2)	(3)	(4)	(5)
2019 sample	9.11**	6.85**	7.45***	2.43	34.68
	(2.95)	(2.69)	(2.60)	(2.49)	(22.62)
Female		-4.38	-5.32*	-5.39**	-1.97
		(3.10)	(3.06)	(2.68)	(2.07)
Age		-0.09	-0.08	-0.02	0.05
		(0.07)	(0.07)	(0.07)	(0.07)
No. of adults		0.57	0.19	0.63	2.72
		(1.89)	(1.89)	(1.68)	(2.47)
No. of children		-3.13**	-2.89**	-2.07*	0.03
		(1.19)	(1.16)	(1.15)	(1.06)
High school		-23.15**	-21.09**	-21.07**	-7.40*
		(10.37)	(10.29)	(9.65)	(3.91)
Post-high school		-21.78**	-20.26**	-20.61**	-3.72
		(9.87)	(9.87)	(9.25)	(4.06)
University graduate		-17.27	-16.50	-18.66*	1.68
		(11.56)	(11.40)	(10.83)	(4.47)
Income		0.15***	0.15***	0.14***	0.03
		(0.03)	(0.03)	(0.03)	(0.03)
No response income		17.12	19.27	17.19	20.79*
		(13.54)	(13.70)	(11.98)	(11.04)
Democrat		(16.35***	5.79	9.28***
			(3.96)	(4.34)	(2.54)
Other party			-0.22	-1.61	5.36
cance parcy			(4.55)	(4.09)	(4.01)
Independent			4.06	0.23	5.68**
macponacine			(3.86)	(3.75)	(2.50)
Temp. has not increased			(0.00)	-4.74**	-4.56*
romp. Had not morododa				(1.61)	(2.35)
Own country should lead				4.99**	0.81
own country should load				(2.07)	(2.36)
Mitigate cost-efficient				5.04	5.58*
Willigate cost efficient				(4.20)	(3.28)
Supp. strict. env. pol.				19.14***	14.49***
oupp. strict. env. poi.				(2.68)	(2.66)
2019 × Female				(2.00)	-6.10
2019 ~ 1 emale					(4.35)
2019 × Age					-0.10
2019 ~ Age					(0.13)
2019 × Adults					-3.56
2019 ^ Adults					
2019 × Children					(3.25) -3.76*
2019 × Children					
0010 LII-k -					(2.20)
2019 × High s.					-24.83*
2010 v Doot hi					(14.80)
2019 × Post-hi.					-27.72**
0010					(13.13)
2019 × Univ.					-31.73**
					(14.75)
2019 × Income					0.18***

					(0.05)
2019 × Demo.					-5.56
					(7.03)
2019 × Other					-10.10
					(6.43)
2019 × Inde.					-8.43
					(5.52)
2019 × Temp. not inc.					0.91
					(3.58)
2019 × Own country					7.13*
					(4.30)
2019 × Mit. cost-efficient					-2.60
					(6.30)
2019 × Stricter env. pol.					8.32*
2010 · Othlotor onto pos					(5.05)
Locational controls	No	Yes	Yes	Yes	Yes
No. observations	2,458	2,458	2,458	2,458	2,458

Note: Dependent variable is WTP for 30% reduction of emissions. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table A6. Sweden: Marginal Effects from Interval Regression Model

	(1)	(2)	(3)	(4)	(5)
2019 sample	4.44***	4.20***	5.13***	6.55***	-14.31*
	(1.25)	(1.23)	(1.25)	(1.24)	(8.67)
Female		-3.29***	-3.64***	-3.71***	-2.13
		(1.24)	(1.23)	(1.22)	(1.61)
Age		-0.12***	-0.12***	-0.12***	-0.20***
		(0.04)	(0.04)	(0.04)	(0.07)
No. of adults		-0.35	-0.19	0.30	-0.82
		(1.04)	(1.00)	(0.93)	(1.47)
No. of children		-1.13	-1.08	-0.82	-1.20
		(0.75)	(0.75)	(0.72)	(1.06)
High school		4.71**	5.36***	4.59**	3.20
		(1.89)	(1.90)	(1.84)	(2.39)
Post-high school		3.92*	4.72**	3.56*	1.71
		(2.13)	(2.14)	(2.08)	(2.62)
University graduate		14.12***	14.65***	12.03***	10.37***
		(2.24)	(2.25)	(2.18)	(2.88)
Income		0.30***	0.30***	0.29***	0.29***
		(0.04)	(0.04)	(0.04)	(0.05)
Centrist			7.98***	4.56**	2.82
			(2.10)	(2.05)	(2.99)
Left			11.43***	6.95***	3.30*
			(1.51)	(1.47)	(1.90)
Other party			-0.45	1.31	-2.50
			(1.86)	(1.83)	(2.31)
Temp. has not increased				-5.84***	-8.01***
				(1.49)	(1.76)
Own country should lead				9.64***	6.51***

				(1.27)	(1.80)
Mitigate cost-efficient				4.20***	3.27**
				(1.26)	(1.64)
Supp. strict. env. pol.				11.84***	10.24***
				(1.19)	(1.50)
2019 × Female					-2.82
0010 4 = -					(2.37)
2019 × Age					0.12
0010 A dulta					(0.08) 1.84
2019 × Adults					(1.93)
2019 × Children					0.62
2019 × Children					(1.44)
2019 × High s.					2.36
2019 ^ Flight 5.					(3.74)
2019 × Post-hi.					2.96
2010 1 1 000 111.					(4.14)
2019 × Univ.					2.57
					(4.18)
2019 × Income					-0.001
					(0.08)
2019 × Centrist					2.22
					(4.01)
2019 × Left					6.25**
					(3.02)
2019 × Other					6.01
					(3.89)
2019 × Temp. not inc.					4.49
					(3.27)
2019 × Own country					5.33*
					(2.73)
2019 × Mit. cost-efficient					2.05
					(2.49)
2019 × Stricter env. pol.					2.84
					(2.46)
Locational controls	No	Yes	Yes	Yes	Yes
No. observations	2,726	2,726	2,726	2,726	2,726

Note: Dependent variable is WTP for 30% reduction of emissions. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table A7. China: Marginal Effects from Interval Regression Model

	(1)	(2)	(3)	(4)	(5)
2019 sample	13.14***	9.02***	9.03***	8.31***	30.94***
	(0.91)	(1.02)	(1.02)	(1.02)	(4.37)
Female		-1.05	-0.94	-1.00	0.55
		(1.07)	(1.06)	(1.06)	(0.46)
Age		-0.47***	-0.46***	-0.47***	0.01
		(0.04)	(0.04)	(0.04)	(0.02)
No. of adults		0.44	0.45	0.52	-0.11

		(0.54)	(0.54)	(0.53)	(0.17)
No. of children		1.80	1.87	1.77	0.36
		(1.14)	(1.14)	(1.137)	(0.32)
High school		6.06***	6.07***	5.37***	-0.71
		(1.17)	(1.17)	(1.19)	(0.48)
Post-high school		8.30***	7.98***	7.03***	-0.67
		(1.15)	(1.17)	(1.188)	(0.64)
University graduate		13.98***	13.10***	12.16***	0.08
		(1.52)	(1.59)	(1.59)	(0.91)
Income		0.06	0.053	0.05	0.35***
		(0.04)	(0.04)	(0.04)	(0.04)
Party member			3.68**	3.08**	0.06
			(1.43)	(1.43)	(0.73)
Temp. has not increased				2.28	-0.36
				(3.41)	(0.47)
Own country should lead				3.03***	1.46***
				(1.11)	(0.44)
Mitigate cost-efficient				2.91**	0.21
				(1.17)	(0.50)
Supp. strict. env. pol.				3.75***	0.21
0010 5				(1.12)	(0.45)
2019 × Female					-1.89 (1.72)
2010 ** 4.75					(1.72)
2019 × Age					-0.71***
2019 × Adults					(0.07) 0.72
2019 ~ Adults					(0.93)
2019 × Children					1.47
2010 · Offination					(1.73)
2019 × High s.					2.33
20.0 × 1.11g.1 0.					(1.90)
2019 × Post-hi.					1.83
					(2.86)
2019 × Univ.					5.66
					(3.64)
2019 × Income					-0.30***
					(0.07)
2019 × Party mem.					7.50***
					(2.40)
2019 × Temp. not inc.					7.69
					(7.76)
2019 × Own country					2.17
					(2.38)
2019 × Mit. cost-efficient					3.02
					(1.89)
2010 v Strictor ony nol					5.96***
2019 × Stricter env. pol.					(2.26)
Locational controls	No	Yes	Yes	Yes	Yes
No. observations	2,843	2,843	2,843	2,843	2,843

Note: Dependent variable is WTP for 30% reduction of emissions. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table A8. United States: Marginal Effects from Interval Regression Model

	(1)	(2)	(3)	(4)	(1)	(2)	(3)	(4)
	60% red	uction comp	ared with 309	% reduction		ction comp	ared with 60	% reduction
2019 sample	8.94***	6.81***	7.20***	3.23**	8.49***	6.62***	6.91***	3.54***
	(2.27)	(1.88)	(1.83)	(1.61)	(2.10)	(1.54)	(1.49)	(1.33)
Female		-4.84*	-5.47*	-5.38**		-3.94	-4.43*	-4.27*
		(2.77)	(2.82)	(2.45)		(2.59)	(2.65)	(2.31)
Age		-0.05	-0.04	0.003		-0.08	-0.08	-0.04
		(0.06)	(0.06)	(0.06)		(0.06)	(0.05)	(0.06)
No. of adults		1.01	0.76	1.11		1.06	0.88	1.16
		(1.71)	(1.66)	(1.48)		(1.73)	(1.72)	(1.54)
No. of		-1.09	-0.92	-0.26		-0.83	-0.71	-0.15
children		(0.76)	(0.78)	(0.76)		(0.67)	(0.67)	(0.63)
High school		-13.90*	-12.46	-12.21*		-11.58	-10.47	-10.11
		(7.89)	(8.12)	(7.23)		(7.35)	(7.65)	(6.78)
Post-high		-11.72	-10.62	-10.70		-10.48	-9.64	-9.59
school		(7.90)	(8.17)	(7.35)		(7.29)	(7.62)	(6.82)
University		-8.07	-7.53	-9.01		-9.13	-8.70	-9.79
graduate		(8.92)	(9.03)	(8.25)		(8.17)	(8.32)	(7.59)
Income		0.10***	0.10***	0.09***		0.10***	0.10***	0.09***
		(0.02)	(0.02)	(0.02)		(0.02)	(0.02)	(0.02)
No response		18.17	19.66	17.90		18.02	19.15	17.58
income		(13.81)	(13.87)	(12.22)		(13.87)	(13.93)	(12.36)
Democrat			10.74***	2.560			7.88**	1.131
			(3.46)	(4.05)			(3.17)	(3.91)
Other party			-0.09	-1.10				-0.78
			(4.16)	(3.76)				(3.77)
Independent			2.67	-0.33			1.55	-0.95
			(3.57)	(3.61)			(3.42)	(3.52)
Temp. has				-3.47***				-2.30**
not increased				(1.26)				(1.06)
Own country				3.78**				3.29**
should lead				(1.78)				(1.62)
Mitigate				5.77				6.00*
cost-efficient				(3.53)				(3.39)
Supp. strict.				14.57***				11.94***
env. pol.				(2.17)				(2.11)
Locational controls	No	Yes	Yes	Yes	No	Yes	Yes	Yes
No.	2,458	2,458	2,458	2,458	2,458	2,458	2,458	2,458
observations	' ' '							,

Note: Dependent variable is WTP for 60% and 85% reduction of emissions compared with 30% and 60%, respectively. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01

Table A9. Sweden: Marginal Effects from Interval Regression Model

	(1)	(2)	(3)	(4)	(1)	(2)	(3)	(4)
	60% redu	ıction compa	red with 30%	reduction	85% redu	iction compa	red with 60%	6 reduction
2019 sample	3.89***	3.84***	4.49***	5.49***	2.93***	2.91***	3.59***	4.32***
	(1.14)	(1.114)	(1.12)	(1.11)	(1.06)	(1.03)	(1.03)	(1.01)
Female		-4.42***	-4.64***	-4.62***		-3.95***	-4.22***	-3.87***
		(1.12)	(1.12)	(1.11)		(1.08)	(1.07)	(1.03)
Age		-0.20***	-0.20***	-0.20***		-0.24***	-0.24***	-0.24***
		(0.04)	(0.04)	(0.04)		(0.04)	(0.04)	(0.04)
No. of adults		-0.03	0.10	0.47		0.31	0.41	0.72
		(1.45)	(1.41)	(1.35)		(1.94)	(1.90)	(1.84)
No. of		-1.92***	-1.89***	-1.70**		-1.74***	-1.69***	-1.53**
children		(0.68)	(0.68)	(0.67)		(0.64)	(0.64)	(0.628)
High school		1.92	2.34	1.75		1.14	1.69	1.19
		(1.79)	(1.84)	(1.82)		(1.75)	(1.79)	(1.77)
Post-high		2.38	2.91	2.06		1.03	1.70	0.97
school		(2.02)	(2.07)	(2.04)		(1.96)	(2.01)	(2.00)
University		7.93***	8.25***	6.30***		6.51***	7.12***	5.46***
graduate		(2.07)	(2.11)	(2.06)		(2.09)	(2.12)	(2.09)
Income		0.24***	0.24***	0.23***		0.15***	0.15***	0.15***
		(0.04)	(0.04)	(0.04)		(0.05)	(0.04)	(0.05)
Centrist			5.67***	3.16			3.63**	1.82
			(1.98)	(1.94)			(1.50)	(1.45)
Left			8.25***	4.93***			8.22***	5.81***
			(1.39)	(1.34)			(1.34)	(1.27)
Other party			-1.32	-0.01			0.51	1.74
			(1.65)	(1.62)			(1.72)	(1.72)
Temp. has				-5.01***				-3.17**
not increased				(1.30)				(1.52)
Own country				6.30***				3.14**
should lead				(1.40)				(1.44)
Mitigate				2.85**				2.54**
efficient				(1.15)				(1.07)
Supp. strict.				9.28***				8.77***
env. pol.				(1.17)				(1.16)
Locational controls	No	Yes	Yes	Yes	No	Yes	Yes	Yes
No. observations	2,726	2,726	2,726	2,726	2,726	2,726	2,726	2,726

Note: Dependent variable is WTP for 60% and 85% reduction of emissions compared with 30% and 60%, respectively. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table A10. China: Marginal Effects from Interval Regression Model

	(1)	(2)	(3)	(4)	(1)	(2)	(3)	(4)
	60% red	uction compa	red with 30%	reduction	85% redu	iction compa	red with 60%	reduction
2019 sample	15.73***	11.99***	12.00***	11.38***	16.73***	13.92***	13.94***	13.33***
	(0.93)	(1.02)	(1.02)	(1.02)	(0.99)	(1.20)	(1.19)	(1.18)
Female		-1.36	-1.19	-1.24		-1.06	-0.84	-0.88
		(1.0)	(1.06)	(1.04)		(1.14)	(1.130)	(1.12)
Age		-0.47***	-0.47***	-0.48***		-0.45***	-0.44***	-0.45***
		(0.04)	(0.04)	(0.04)		(0.04)	(0.04)	(0.04)
No. of adults		-0.02	-0.01	0.05		0.004	0.02	0.08
		(0.53)	(0.53)	(0.53)		(0.59)	(0.58)	(0.58)
No. of kids		0.55	0.65	0.57		0.58	0.71	0.63
		(0.76)	(0.75)	(0.75)		(0.81)	(0.80)	(0.80)
High school		6.58***	6.61***	6.04***		6.65***	6.69***	6.11***
		(1.18)	(1.18)	(1.196)		(1.28)	(1.28)	(1.29)
Post-high		8.21***	7.72***	6.91***		9.35***	8.74***	7.93***
school		(1.19)	(1.21)	(1.23)		(1.28)	(1.28)	(1.31)
University		13.44***	12.07***	11.27***		14.39***	12.70***	11.89***
graduate		(1.53)	(1.59)	(1.60)		(1.67)	(1.67)	(1.68)
Income		0.06	0.05	0.05		0.02	0.002	0.001
		(0.04)	(0.04)	(0.04)		(0.05)	(0.05)	(0.05)
Party			5.82***	5.27***			7.25***	6.71***
member			(1.687)	(1.68)			(1.85)	(1.85)
Temp. has				3.36				2.61
not								
increased				(3.44)				(2.67)
Own country				2.75***				2.78***
should lead				(1.02)				(1.03)
Mitigate cost-				2.52**				2.06
efficient				(1.17)				(1.26)
Supp. strict.				3.55***				3.46***
env. pol.				(1.02)				(1.07)
Locational controls	No	Yes	Yes	Yes	No	Yes	Yes	Yes
No. observations	2,843	2,843	2.843	2.843	2,843	2,843	2,843	2,843

Note: Dependent variable is WTP for 60% and 85% reduction of emissions compared with 30% and 60%, respectively. Standard errors in parentheses. * p < 0.1, ** p < 0.05, *** p < 0.01.

Table A11. Monthly Household WTP in PPP Dollars (Survey Weights) among Democrats, Republicans, and Independents

	Democrats		Democrats Republicans			ans	s Independents	
Reduction level	Mean	St. dev.	Mean	St. dev.	Mean	St. dev.		
2019								
30% reduction	45	59	15	35	25	45		
60% reduction	77	95	29	69	47	79		
85% reduction	103	129	40	103	65	111		
Share zero WTP	0.05		0.37		0.18			
		200	9					
30% reduction	28	48	10	24	20	43		
60% reduction	47	81	15	32	33	70		
85% reduction	61	108	18	37	42	91		
Share zero WTP	0.12		0.35		0.25			

Table A12. Absolute and Relative Differences in WTP for Three Different Comparisons of Political Preferences in the US Sample between 2019 and 2009

	30% reduction	60% reduction	85% reduction						
Difference in absolute level of WTP									
Democrats - Republicans	10.98	16.12*	20.41						
Democrats - Independents	10.68	15.17	19.25						
Independents - Republicans	0.30	0.95	1.15						
	Difference in relative lev	rels of WTP							
Democrats / Republicans	0.00	-0.46	-0.75						
Democrats / Independents	0.32	0.19	0.14						
Independents / Republicans	-0.37	-0.53	-0.67						

Note: P-values for the z-test were retrieved using the Krinsky and Robb (1986) method. Survey weights based on the population level of educational attainment are applied for both political parties and independents in the United States in both 2009 and 2019 (survey weights are available on request). We used 1,000 draws from a normal distribution with mean and variances given by the estimated mean and variance of WTP. p < 0.1, p < 0.05, p < 0.01.

Table A13. Attitudes toward Climate Change and Policy among Democrats, Republicans, and Independents in the United States

			Democr	ats		Republicans			Independents	
		2009	2019	p-val	2009	2019	p-val	2009	2019	p-val
No temperature increase	The temperature has not increased globally.	10%	5%	0.001	40%	32%	0.008	28%	17%	0.058
Humans affect temperature increase	Humans have affected the temperature increase.	86%	97%	0.000	57%	68%	0.008	72%	79%	0.240
Cannot stop climate change	We cannot do anything to stop climate change.	14%	7%	0.030	22%	21%	0.935	16%	13%	0.354
Can mitigate but not stop	We can mitigate but not stop climate change.	68%	73%	0.246	51%	55%	0.327	59%	66%	0.194
Can stop climate change	We can stop climate change.	14%	17%	0.317	6%	7%	0.680	9%	15%	0.200
Climate denial	The climate is not changing.	3%	1%	0.005	21%	17%	0.229	16%	4%	0.000
Own country should reduce	Own country should reduce carbon emissions, even if other countries do not reduce their carbon emissions.	80%	96%	0.000	58%	61%	0.404	68%	74%	0.342
Reduce where it is cheap	Own country should primarily use public funding to reduce carbon emissions wherever it is cheapest, even if it means in another country.	18%	32%	0.001	14%	17%	0.416	17%	29%	0.032
Reduce in own country	Own country should primarily use public funding to reduce carbon emissions in own country.	51%	54%	0.488	37%	45%	0.058	43%	42%	0.818
No opinion	No opinion or other opinion.	32%	14%	0.000	49%	38%	0.014	40%	30%	0.072
Primarily mitigate	Countries in the world should deal with climate change primarily by reducing carbon emissions.	74%	88%	0.000	42%	50%	0.069	58%	60%	0.812
Primarily adaptation	Countries in the world should deal with climate change primarily by adaptation.	13%	10%	0.531	16%	20%	0.176	18%	25%	0.221
Do nothing	Do nothing, because I don't consider climate change to be a problem.	14%	2%	0.000	42%	30%	0.003	24%	15%	0.009
Trustworthy information	The information given in the survey regarding climate change is trustworthy.	67%	87%	0.000	32%	41%	0.037	49%	57%	0.142
Prioritize environment	We should prioritize environmental improvements, even if we lose jobs.	52%	88%	0.000	30%	37%	0.069	39%	55%	0.005
No. observations		368	500		248	390		284	436	

Note: P-values for t-test of difference in means based on survey weights.

Table A14. Monthly Household WTP in PPP Dollars (Survey Weights among Swedish Left-Wing Voters, Right-Wing Voters, and Centrist Voters)

	Left-wing	voters	Right-wing voters		Centrist	voters			
Reduction level	Mean	St. dev.	Mean	St. dev.	Mean	St. dev.			
2019									
30% reduction	38	48	21	35	33	43			
60% reduction	68	88	40	66	60	79			
85% reduction	94	127	54	92	79	104			
Share zero WTP	0.04	0.04		0.18					
		200	9						
30% reduction	26	36	21	34	26	35			
60% reduction	49	67	37	59	48	66			
85% reduction	67	99	48	83	63	90			
Share zero WTP	0.04		0.12		0.06				

Table A15. Absolute and Relative Differences in WTP for Three Different Comparisons of Political Preferences in the Swedish Sample between 2019 and 2009

	30% reduction	60% reduction	85% reduction					
Difference in absolute level of WTP								
Left - Right	11.6***	16.4**	21.6**					
Left - Centrist	4.5	7.8	11.2					
Centrist - Right	7.2	8.6	10.4					
	Difference in relative le	vels of WTP						
Left / Right	0.55**	0.38*	0.36					
Left / Centrist	0.13	0.13	0.13					
Centrist / Right	0.34	0.19	0.16					

Note: P-values for the z-test were retrieved using the Krinsky and Robb (1986) method. Survey weights based on the population level of educational attainment are applied for all political groups in Sweden in both 2009 and 2019 (survey weights for political parties are available on request). We used 1,000 draws from a normal distribution with mean and variances given by the estimated mean and variance of WTP. \dot{p} < 0.1, \dot{p} < 0.05, \dot{p} < 0.01.

Table A16. Absolute and Relative Differences in WTP for Three Different Comparisons of Political Preferences in the Swedish Sample (excluding the party Sweden Democrats)

	30% reduction	60% reduction	85% reduction						
Difference in absolute level of WTP									
Left - Right	9.1**	10.7	27.2**						
Left - Centrist	4.5	7.7	11.2						
Centrist - Right	4.6	2.9	16.0						
	Difference in relative leve	els of WTP							
Left / Right	0.33	0.18	0.50*						
Left / Centrist	0.13	0.13	0.13						
Centrist / Right	0.16	0.03	0.29						

Note: Excludes the populistic party Sweden Democrats. Survey weights based on the correct level of educational attainment of the population are adjusted after excluding the Swedish Democrats from the comparisons (these weights are available on request). We used 1,000 draws from a normal distribution with mean and variances given by the estimated mean and variance of WTP. $\dot{p} < 0.1$, $\dot{p} < 0.05$, $\dot{p} < 0.01$.

Table A17. Attitudes toward Climate Change and Policy among Democrats, Republicans, and Independents in Sweden

			Left-wi	ng		Right-wing			Centris	Centrist	
		2009	2019	p-val	2009	2019	p-val	2009	2019	p-val	
No temperature increase	The temperature has not increased globally.	5%	2%	0.037	7%	10%	0.146	3%	3%	0.794	
Humans affect temperature increase	Humans have affected the temperature increase.	96%	98%	0.024	92%	93%	0.420	99%	98%	0.471	
Cannot stop climate change	We cannot do anything to stop climate change.	6%	4%	0.164	8%	7%	0.610	1%	6%	0.048	
Can mitigate but not stop	We can mitigate but not stop climate change.	80%	79%	0.804	82%	77%	0.120	83%	84%	0.723	
Can stop climate change	We can stop climate change.	13%	17%	0.119	6%	11%	0.014	15%	9%	0.109	
Climate denial	The climate is not changing.	2%	1%	0.006	3%	2%	0.474	0%	1%	0.318	
Own country should reduce	Own country should reduce carbon emissions, even if other countries do not reduce their carbon emissions.	89%	89%	0.712	73%	60%	0.001	79%	83%	0.399	
Reduce where it is cheap	Own country should primarily use public funding to reduce carbon emissions wherever it is cheapest, even if it means in another country.	38%	36%	0.559	46%	50%	0.273	60%	39%	0.000	
Reduce in own country	Own country should primarily use public funding to reduce carbon emissions in own country.	48%	47%	0.703	32%	27%	0.148	32%	42%	0.086	
No opinion		13%	16%	0.191	22%	23%	0.759	8%	19%	0.010	
Primarily mitigate	No opinion or other opinion.	93%	92%	0.727	75%	74%	0.823	91%	88%	0.393	
Primarily adaptation	Countries in the world should deal with climate change primarily by reducing carbon emissions.	5%	6%	0.554	18%	20%	0.573	6%	12%	0.180	
Do nothing	Countries in the world should deal with climate change primarily by adaptation.	2%	1%	0.729	7%	6%	0.571	2%	1%	0.373	
Trustworthy information	Do nothing, because I don't consider climate change to be a problem.	86%	87%	0.561	77%	71%	0.079	87%	88%	0.764	
Prioritize environment	The information given in the survey regarding climate change is trustworthy.	68%	70%	0.496	60%	42%	0.000	57%	67%	0.110	
# Observations	We should prioritize environmental improvements, even if we lose jobs.	532	525		404	603		161	215		

Note: P-values for t-test of difference in means based on survey weights.

Online Appendix. Assumptions and conversion of monthly household willingness to pay (WTP) to WTP/ton CO₂

To be able to compare the willingness to pay estimates from our study to the social cost of carbon estimates most frequently referenced in the literature, we convert the monthly household willingness to pay to WTP per ton of carbon emissions (dollars/ton CO_2).

Definitions

 W_{CY}^H = WTP per month for a household in country C and year Y

 H_{CY} = Number of households in country C in year Y

 E_{CY} = emissions (Mt CO₂) in country C in year Y

 W_{CY}^{SCC} = WTP in country C and year Y expressed in dollars/ton CO₂

The WTP per ton of carbon emissions (dollars/ton CO₂) in country C in year Y is calculated as the total (annual) estimated willingness to pay divided by the reduction in emissions.

$$W_{CY}^{SCC} = \frac{W_{CY}^H * H_{CY} * 12}{0.3 * E_{CY}}$$

Note that the reduction in emissions is calculated as the percentage reduction (30 percent) of total emissions for the base year Y (Table O1). The results for the three countries and years are presented in Table A1; monthly WTP values for a 30 percent reduction are obtained from Table 3.

Table O1. WTP in Dollar per Ton CO₂ (at a 30% reduction level)

	2009	2019
United States	18	31
Sweden	85	129
China	11	44

Data Sources

Emissions data from Global Carbon Atlas

Gilfillan, D., G. Marland, T. Boden, and R. Andres, 2019. "Global, Regional, and National Fossil-Fuel CO₂ Emissions." Boone, NC: Carbon Dioxide Information Analysis Center at Appalachian State University. https://energy.appstate.edu/research/work-areas/cdiac-appstate (accessed September 27, 2019).

UNFCCC (United Nations Framework Convention on Climate Change). 2019. "National Inventory Submissions 2019." https://unfccc.int/process-and-meetings/transparency-and-reporting/reporting-and-review-under-the-convention/greenhouse-gas-inventories-annex-i-parties/national-inventory-submissions-2019 (accessed June 11, 2019).

BP. 2019. "Statistical Review of World Energy." https://www.bp.com/en/global/corporate/energy-economics/statistical-review-of-world-energy.html (accessed June 11, 2019).

Population data

World Bank. https://data.worldbank.org/country/ (accessed November 16, 2020).

Number of persons per household

Sweden: Data from Statistics Sweden 2019 and 2012. https://www.statistikdatabasen.scb.se/ (accessed November 16, 2020).

United States: Data from https://www.statista.com/statistics/183648/average-size-of-households-in-the-us/ (accessed November 16, 2020).

China: Data from https://www.statista.com/statistics/278697/average-size-of-households-in-china/ (accessed November 16, 2020).